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## DISPERSION - A NEW APPROACH

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Dispersion of a set of observations $x_{1}, x_{2}, x_{3}, \ldots \ldots . x_{n}$ is measured by variance, $s^{2}=(1 / n-1) \sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)^{2}$

In this paper it is shown that $s^{2}$ can be expressed in terms of the sum of squares of differences between all possible pairs of observations. The implications of this new definition are studied in a number of situations.

## Results

The expression for variance $s^{2}=(1 / n-1) \sum_{i=1}^{n}\left(x_{i}-x\right)^{2}$ can be simplified to the form,
$s^{2}=\left\{\frac{1}{n(n-1)}\right\} \quad \sum_{i}>\sum_{j}\left(x_{i}-x_{j}\right)^{2}$
Proof

$$
\begin{aligned}
s^{2} & =\frac{1}{(n-1)}-\sum_{i}(v,-x)^{2} \\
& =\frac{1}{(n-1)}\left\{\sum_{i} x_{i}^{2} \quad \sum_{i} x_{1}^{2},-\sum \sum_{i} \sum_{j} x_{1} x_{i}\right\} \\
& =\frac{1}{n(n-1)} \sum_{i>j} \sum_{j}\left(x_{i}-x_{j}\right)^{2}
\end{aligned}
$$

Thus $s^{2}$ is $-\frac{1}{n(n-1)}$ - times the sum of squares between pairs of observations $x_{i}, x_{j}$ such that $i>j$. Further, this can be partitioned into ( $n-1$ ) components $d_{i}^{2}, d_{c}^{a}, \ldots . \quad d_{n-1}^{2}$, where $d^{2} i=\frac{n-j}{i}=\sum_{1}^{j}\left(x_{i} ;-x_{i}\right)^{2}$, is the. sum of squares of differences between pairs of observations which are j units apart. Thus $d_{1}^{a}$ is the sum of squares $\sum_{i=1}^{n-1}\left(x_{i, 1}, 1 \cdot x_{i}\right)^{2}$. which on division by $(n-1)$ gives the mean square successive difference $D_{i}^{2}$. Thus $D_{1}^{2}=\frac{d_{1}^{2}}{(n-1)} \quad$ Similarly $D_{2}^{2}, D_{3}^{2} \ldots D_{11, \ldots}^{2}$, can be defined as the mean square differences of orders
$2,3, \ldots(n-1)$, respectively, where
$D_{2}^{2}=\frac{d_{L}^{2}}{(n-2)}, v_{3}^{?}=\frac{d_{3}^{2}}{(n-3)}, \ldots \ldots D_{n-\hat{i}}^{2}=d_{n-1}^{2}$
These are different from the mean square successive $j$-th difference $\delta_{\text {, }}^{2}$ where

$$
\delta_{1}^{2}=\frac{1}{(n-1)} \sum_{i=1}^{n-1}\left(x_{i+1}-x_{i}\right)^{2} \cdot D_{1}^{2}
$$

$$
{\underset{i}{\mathrm{t}}}_{2}=\sim_{i=1}^{n-2} \sum_{i=1}^{n-<1}\left(x_{i+<}-2 x_{i+1}+x_{i}\right)^{2}
$$

$$
\delta_{3}^{2}=\frac{1}{(n-3)} \sum_{i=1}^{n-3}\left(x_{i+3}-3 x_{i+2}+3 x_{i+1}-x_{i}\right)^{2}
$$

$$
\delta_{4}^{2}=\frac{1}{(n-4)} \sum_{i=1}^{n-4}\left(x_{i+4}-4 x_{i+3}+6 x_{i+2}-4 x_{i+1}+x_{i}\right)^{2}
$$

The measure of dispersion, $s^{2}$ can be represented in terms of $D_{1}^{3}$, as shown below:

$$
\begin{aligned}
s^{2} & \ldots \cdot \cdot\left(d_{1}^{2}+d_{2}^{2}+\ldots+{ }^{d^{2}}{ }_{n-1}\right) \\
& =\frac{1}{n(n-1)}\left[(n-1) D_{1}^{2}+(n-2) D_{2}^{2}+\cdots \cdot+D_{n-1}^{2}\right] \\
& =\frac{1}{n(n-1)} \sum_{i=1}^{n-1}(n-i) \quad D_{i}^{2}
\end{aligned}
$$

Hence $s^{2}$ is $(1 / n)$ times the simple average of $d_{i}^{2}$ values or is half of the weighted average of the $D \stackrel{?}{?}$ values.

These lead us to a number of useful results which are stated below :
(a) The sum of squares of $n$ observations $x_{1}, x_{2}, x_{3}, \ldots x_{n}$ namely $\sum\left(x_{i}-\bar{x}\right)^{2}$ can be expressed as
(if)

$$
\begin{aligned}
\sum_{i>j}\left(x_{i}-x_{j}\right)^{2}= & \sum_{i=1}^{n-1}\left(x_{i+1}-x_{i}\right)^{2}
\end{aligned}+\sum_{i+1}^{n-2}\left(x_{i+2}-x_{i}\right)^{2}+\ldots \ldots .
$$

(b) The sum of squares between $v$ samplas of sizes $r_{1}, r_{\mathbb{T}} \ldots r_{v}$ having sample totals $T_{1}, T ;, . . . T_{v}$ is given by
$\left(\frac{1}{N}\right) \quad \sum_{i>j}^{<-\pi}\left(r_{i} T_{j}-r_{j} T_{i}\right)^{2} / r_{i} r_{j}$ where $N$ is the sum $r_{1}+r_{2}+\ldots+r_{v}$
Obviously when sample sizes are equal [to $r$, this simplifies to
$\left(\frac{1}{v r}\right) \sum_{i>j} \sum_{j}\left(T_{i}-T_{j}\right)^{2}$ This shows that the $S S$ between samples is derived as the sum of squares of differences between all possible pairs of treatment totals adjusted by dividing by the total number of plots. The contribution of each pair can be easily seen and in fact is indicative of the distances between samples in the one-dimensional space.
(c) The formula for computation of variance in frequency tables can be expressed
as $\frac{1}{\operatorname{ND}(\mathbb{N}-1)} \quad \sum_{i>j}^{>}\left(f_{i} . f_{i}\right)\left(x_{i}-x_{i}\right)^{2}$
Variance is defined as $s^{2}=\frac{1}{N-1} \sum_{i=1}^{N} f_{i}\left(x_{i}-\bar{x}\right)^{\frac{2}{2}}$

When the classes are of uniform width c , variance can be computed by using the formula,
$s^{2}=\underset{N(N-1)}{c^{2}} \sum_{i>j} \sum_{j} f_{i} \cdot f_{j}(i-j)^{2}$
(d) Similar to the simplified form of $\mathrm{s}^{2}$, the covariance between x and y or the product moment can be put in the form,
$P_{x y}=\left[\frac{1}{n-1}\right] \quad \sum_{i=j}\left(x_{i}-x_{j}\right) \quad\left(y_{i}-y_{j}\right)$
(e) In the analysis of variance of data relating to $v$ samples of sizes $r$ each, the SS for total is ${ }^{\prime 1} \sum_{i=j} \sum_{i j}\left(y_{i j}-y_{i}{ }^{j}\right)^{2}$ and $S S$ between samples is $\frac{1}{r v} \sum_{i>j} \sum_{j}\left(y . j^{j}\right)^{2} . j>j$.
(f) In randomised block experiments, block SS will be,
$\frac{1}{r v} \quad i>j\left(y_{i .}-y_{i .}\right)^{2}$
(g) The new method of defining the $\mathbf{S S}$ makes the method of estimation of missing plot yields simple.

Consider a randomised block experiment with $v$ treatments and $r$ replications. Where $y_{i j}$ is missing the method of estimation of $y_{i j}$ is by minimising error SS $\left.\frac{1}{r v}\left[\sum_{i^{\prime}} \sum_{j^{\prime}} y_{i j}-y_{i^{\prime} j}\right)^{3}-\sum_{i^{\prime}}\left(y_{i}-y_{i^{\prime}}\right)^{2}-\sum_{j^{j}}\left(y_{j}-y_{j^{\prime}}\right)^{2}\right]$
This leads to the equation,

$$
\begin{aligned}
& \sum_{i^{\prime} j^{\prime}}\left(y_{i i}-y_{i^{\prime} j^{\prime}}\right)-\sum_{i^{\prime}}\left(y_{i j}-y_{i^{*}}\right)+\sum_{j^{\prime}}\left(y_{j}-y_{\cdot j}\right) \\
& (r v-1) y_{i j}-G^{\prime}=(r-1) y_{i}-\left(G^{\prime}-y_{i}\right)+(v-1) y_{j} \cdot-\left(G^{\prime}-y\right) \\
& =r \cdot y_{i}-G^{\prime}+v \cdot y_{i}-G^{\prime}
\end{aligned}
$$

Where $\mathrm{G}^{\prime}=$ total of $\mathrm{y}^{\prime} \mathrm{s}$ except from the missing plot.
Thus the estimate of $y_{i j}$ is $\frac{r \cdot y:+v \cdot y_{i}^{\prime}-G^{\prime}}{(r-1)(v-1)}$

## Summary

The variance $s^{2}$ can be expressed in terms of the sum of squares of differences between all possible pairs of observations. This new approach to the definition of $s^{2}$ is useful in understanding many concepts in statistics and also provides simpler computational formulae. Implications of this new definition are studied in a number of situations.

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