

PROPERTIES OF THE SAMPLE VARIANCE IN SAMPLES FROM A GENERALISED AUTOREGRESSIVE SCHEME

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The distribution of sample variance $s^2 = 2(x_i - \bar{x})^2 / (n-1)$, where the x_i 's are randomly and independently drawn from a normal population with mean μ and variance S^2 , is,

$$P(s^2) = \frac{\left(\frac{n-1}{2S^2}\right)^{\frac{n-1}{2}}}{\Gamma\left(\frac{n-1}{2}\right)} \cdot e^{-\frac{(n-1)s^2}{2S^2}} \cdot \frac{1}{2} \cdot (s^2)^{\frac{n-1}{2}-1}$$

(Rietz, 1937J. The moments of s^2 are as follows. Mean $(M_1') = S^2$, variance $(M_2) = \left(\frac{2}{n-1}\right) \cdot S^4$, $M_3 = \frac{8}{(n-1)^2} \cdot S^6$, $M_4 = \frac{12(n+3)}{(n-1)^3} \cdot S^8$. Thus $\beta_1 = 8/(n-1)$ and $\beta_2 = 3 + \frac{12}{n-1}$. ('Student', 1908)

In this paper, x_i 's are assumed to follow a generalised autoregressive scheme,

$$X_t + B_1 \cdot X_{t-1} + B_2 \cdot X_{t-2} + \dots + B_p \cdot X_{t-p} = e_t + C_1 \cdot e_{t-1} + C_2 \cdot e_{t-2} \dots + C_n \cdot e_{t-n},$$

where the e_t 's follow a normal distribution with mean 0 and variance S^2 .

Further, e_t 's are assumed to be independent. Under this scheme, the properties of s^2 are investigated. The algebra involved in the computation of the third and fourth moments tend to be laborious. The methods are utilised to find the first four moments and also β_1 and β_2 , for the simpler model, $x_t = e_t + a \cdot e_{t-1}$.

Materials and Methods

The variable x_t follows the generalised autoregressive scheme,

$$x_t + B_1 \cdot x_{t-1} + B_2 \cdot x_{t-2} + \dots + B_p \cdot x_{t-p} = e_t + C_1 \cdot e_{t-1} + C_2 \cdot e_{t-2} + \dots + C_n \cdot e_{t-n}.$$

This can be written in the form,

$$x_t = e_t + a_1 \cdot e_{t-1} + a_2 \cdot e_{t-2} + \dots \quad \text{where,}$$

$$\frac{1 + C_1 \cdot u + C_2 \cdot u^2 + \dots + C_n \cdot u^n}{1 + B_1 \cdot u + B_2 \cdot u^2 + \dots + B_p \cdot u^p} = 1 + a_1 \cdot u + a_2 \cdot u^2 + \dots$$

Since the e, 's are assumed to be normally and independently disiributed with mean 0 and variance S^2 , it follows that, $E(x_t) = 0$.

$$E(x_t^2) = (1 + a_1^2 + a_2^2 + \dots). S^2 = A_0 \cdot S^2 \text{ (say) and}$$

$$E(x_t \cdot x_{t+k}) = (a_k + a_1 \cdot a_{k-1} + a_2 \cdot a_{k-2} + \dots) S^2 = A_k \cdot S^2, \text{ where}$$

$$A_k = a_k + a_1 \cdot a_{k+1} + a_2 \cdot a_{k+2} + \dots$$

To obtain the four moments of s^2 , the usual procedure of finding expectations of the first four powers of s^2 is adopted. This involves expansions of the from $(\sum x_t^2)^p \cdot C \sum x_t^2)^{2q}$, where p and q are zero or positive integers such that $p < 4$, $q < 4$, and $p + q = 4$.

$$(\sum x_t^2)^2 = \sum x_t^4 + 2 \sum_{r=1}^{n-1} \sum_{t=1}^{n-r} x_t^2 \cdot x_{t+r}^2$$

$$(\sum x_t^2)^3 = \sum x_t^6 + 3 \sum_{r=1}^{n-1} \sum_{t=1}^{n-r} S x_t^4 \cdot x_{t+r}^2 + 6 \sum_{r=1}^{n-2} \sum_{s=1}^{n-r-s} S$$

$$\sum_{t=1}^{n-r-s} x_t^2 \cdot x_{t+r}^2 \cdot x_{t+r+s}^2$$

where $\sum_{t=1}^{n-r-s} x_t^4 \cdot x_{t+r}^2 = \sum_{t=1}^{n-r-s} x_t^2 \cdot x_{t+r}^2 \cdot x_{t+r+s}^2$

The other expressions, being still more lengthy, are not listed here,

Expectations of terms like $x_t^{2k_1} \cdot x_{t+r}^{k_2} \cdot x_{t+r+s}^{k_3} \cdot x_{t+r+s+w}^{k_4} \cdot x_{t+r+s+w+v}^{k_5}$

if $x_t \cdot x_{t+s} \cdot x_{t+w} \cdot x_{t+u}$, where k_2, k_3, \dots, k_7 can take values from 0 to 7 and k_1 takes values from 0 to 4, such that $2 \cdot k_1 + k_2 + k_3 + k_4 + k_5 + k_6 + k_7 < 8$, are required for the first four moments of s^2 . The following results are easily derived.

$$E(x_t) = 0, E(x_t^2) = A_0 \cdot S^2, E(x_t \cdot x_{t+r}) = A_r \cdot S^2, E(x_t^4) = 3 \cdot A_0^2 \cdot S^4$$

$$E(x_t^2 \cdot x_{t+r}^2) = (A_0^2 + 2A_r^2) S^4, \sum_{t=1}^n E(x_{t+r}^3 \cdot x_{t+t}) = 6 \cdot A_0 \cdot A_r \cdot S^4$$

$$E(x_t \cdot x_{t+r} \cdot x_{t+r+s} \cdot x_{t+r+s+w}) = (A_r \cdot A_w + A_s \cdot A_{r+s+w} + A_{r+s} \cdot A_{s+w}) \cdot S^4$$

$$E(x_t^6) = 15 \cdot A_0^3 \cdot S^6, \sum_{t=1}^n E(x_t^5 \cdot x_{t+r}) = 30 \cdot A_0^2 \cdot A_r \cdot S^6$$

$$\sum_{t=1}^n E(x_t^4 \cdot x_{t+r}^2) = (6 \cdot A_0^3 + 24 \cdot A_0 \cdot A_r^2) S^6, E(x_t^5 \cdot x_{t+r}^3) = (9 \cdot A_0^2 \cdot A_t + 6 \cdot A_r^3) \cdot S^6$$

Expectons of other expressions are still more lengthy and hence are not reproduced here.

Results and discussion

Using the methods described, it is possible to derive the moments of s^2 , where $s^2 = \frac{\sum (x_i - \bar{x})^2}{(n-1)} = \frac{(\sum x_i^2)}{(n-1)} - \frac{(\sum x_i)^2}{n(n-1)}$

$$\text{Thus, } M_1' = \left(A_0 - \frac{2}{n(n-1)} \sum_{r=1}^{n-1} (n-r) A_r \right) S^2$$

$$M_2 = \frac{S^4}{n^2 (n-1)^2} \left[2n^2 (n-1) \cdot A_0^2 + 4n (n-2) \sum_{r=1}^{n-1} (n-r) \cdot A_r^2 + 4n (n-4) \sum_{r=1}^{n-1} (n-r) A_0 A_r + 8 \sum_{r=1}^{n-1} \sum_{s=1}^{n-1} (n-r) \cdot A_r^2 \cdot A_s - 4n \sum_{r=1}^{n-1} \sum_{s=1}^{n-1} (n-r-s) A_0 (A_r + A_s + A_{r+s}) + 2 (A_r \cdot A_s + A_r \cdot A_{r+s} + A_s \cdot A_{r+s}) \right]$$

Expressions for M_3 and M_4 in the general case happen to be very lengthy and hence are not given in this paper.

For a simple particular case it is easy to work out the moments. Hence the model, $x_t = e_t + a$. $e_t - 1$ is taken to illustrate the method. In this case, $A_0 = 1 + a^2$, $A_1 = a$ and $A_p = 0$ for p greater than 1. This leads to the following results.

$$M_1' = S^2 [1 - (2a/n) + a^2]$$

$$M_2 = \frac{S^4}{n^2(n-1)^2} \left[2n^2 (n-1) - 8n (n-1) \cdot a + 8 (n^3 - 2n^2 + n + 1) \cdot a^2 - 8n (n-1) \cdot a^3 + 2n^2 (n-1) \cdot a^4 \right]$$

$$M_3 = \frac{S^6}{n^3(n-1)^3} \left[8n^3 (n-1) - 48n^2 (n-1) \cdot a + 24n (3n^3 - 7n^2 + 4n + 4) \cdot a^2 - 32(5n^3 - 6n^2 - 3n - 2) \cdot a^3 + 24n (3n^3 - 7n^2 + 4n + 4) \cdot a^4 - 48n^2 (n-1) \cdot a^5 + 8n^3 (n-1) \cdot a^6 \right]$$

$$M_4 = \frac{S^8}{n^4(n-1)^4} \left[12n^4(n-1)(n+3) - 16n^2(n-1)(6n^2 + 18n - 3681) \cdot a + 96n (n^5 + 2n^4 + 3n^3 - 40n^2 + 106n - 90) \cdot a^2 + 48n (8n^4 - 58n^3 - 3457n^2 + 3329n + 320) \cdot a^3 + 24 (9n^6 + 14n^5 + n^4 - 252n^3 + 684n^2 - 280n + 40) \cdot a^4 + 48n (8n^4 - 58n^3 - 3457n^2 + 3329n + 320) \cdot a^5 + 96n (n^5 + 2n^4 + 3n^3 - 40n^2 + 106n - 90) \cdot a^6 - 8n^2 (n-1) (12n^2 + 36n - 7362) \cdot a^7 + 12n^4 (n-1) (n+3) \cdot a^8 \right]$$

When $a = 0$, these reduce to $M_1' = S^2$, $M_2 = \left(\frac{2}{n-1}\right) S^4$,

$$M_3 = \frac{8}{(n-1)^2} S^6$$

$M_4 = \frac{12(n-1)^2}{(n-1)^3} S^8$, agreeing with the results for the case where x_t, s_t are randomly and independently distributed.

When $a \neq 0$ also, it can be seen that β_1 tends to 0 and β_2 tends to 3 as n tends to infinity. It is difficult to get the exact distribution of s^2 in this case. But the distribution can be approximated by a normal distribution with mean $(1+a^2) \cdot S^2$ and variance $(2/a) \cdot (a^4 + 8n^2 + 2) \cdot S^4$. It is obvious that s^2 is a biased estimate of s^2 for $a \neq 0$

Summary

An attempt is made to study the properties of the sample variance in samples from a generalised autoregressive scheme, by computing the moments. Eventhough the expressions are lengthy, the moments can be derived. For the simple particular case $x_t = e_t + a \cdot e_{t-1}$, the moments have been worked out. The exact distribution has not been derived, but the large sample approximation is found to be normal with mean $S^2 (1 + a^2)$ and variance $(2/n) (a^4 + 8a^2 + 2) \cdot S^4$. When $a \neq 0$, s^2 is a biased of S^2 .

സംഗ്രഹം

സാമാന്യ വൻകൃത സ്വയംസമാശ്രയണ പദ്ധതിയിൽ നിന്നുള്ള സാമ്പിൾ വ്യതിയാനത്തിന്റെ ഗുണധർമ്മങ്ങളെപ്പറ്റി പഠിക്കുവാൻ ഇതിന്റെ ആലുപ്നങ്ങളുടെ പരീക്ഷണം വഴി ഒരു ശ്രമം നടത്തിയിരിക്കുന്നു. വ്യംജകങ്ങൾ വളരെ ദൈർഘ്യമുള്ളവയാണെന്നു വരികിലും ആലുപ്നങ്ങളെ വ്യൽപാദിപ്പിക്കുവാൻ സാധ്യമാണ്. $x_t = e_t + a \cdot e_{t-1}$ എന്ന ലഘു വിശേഷ സ്ഥിതിയിൽ ആലുപ്നങ്ങളെ പരീക്ഷണം ചെയ്തിട്ടുണ്ട്. കൃത്യമായ വിതരണ ഫലനങ്ങൾ വ്യൽപാദിപ്പിക്കപ്പെട്ടിട്ടില്ലെങ്കിലും വിതരണത്തിന്റെ ബൃഹത് സാമ്പിൾ ഏകദേശം, $S^2 (1 + a^2) \frac{2}{n} (a^4 + 8a^2 + 2) S^4$ വ്യത്യസ്തവുമായുള്ള നോർമൽ വിതരണമാണെന്നു കാണുന്നു. $a \neq 0$ ആയിരിക്കുമ്പോൾ s^2 എന്നത് S^2 ന്റെ അഭിനതിയുള്ള ആകലനം ആകുന്നു.

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